d-QPSO: A Quantum-Behaved Particle Swarm Technique for Finding *D*-Optimal Designs With Discrete and Continuous Factors and a Binary Response

Joshua Lukemire^a, Abhyuday Mandal^b, and Weng Kee Wong^c

^aDepartment of Biostatistics and Bioinformatics, Emory University, Atlanta, GA; ^bDepartment of Statistics, University of Georgia, Athens, GA; ^cDepartment of Biostatistics, University of California at Los Angeles, Los Angeles, CA

ABSTRACT

Identifying optimal designs for generalized linear models with a binary response can be a challenging task, especially when there are both discrete and continuous independent factors in the model. Theoretical results rarely exist for such models, and for the handful that do, they usually come with restrictive assumptions. In this article, we propose the *d*-QPSO algorithm, a modified version of quantum-behaved particle swarm optimization, to find a variety of *D*-optimal approximate and exact designs for experiments with discrete and continuous factors and a binary response. We show that the *d*-QPSO algorithm can efficiently find locally *D*-optimal designs even for experiments with a large number of factors and robust pseudo-Bayesian designs when nominal values for the model parameters are not available. Additionally, we investigate robustness properties of the *d*-QPSO algorithm-generated designs to various model assumptions and provide real applications to design a bio-plastics odor removal experiment, an electronic static experiment, and a 10-factor car refueling experiment. Supplementary materials for the article are available online.

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1. Introduction

Our work is motivated by an odor removal study (Wang et al. 2016) conducted in the Department of Textiles, Merchandising, and Interiors within the College of Family and Consumer Sciences at the University of Georgia (UGA). In the research, protein-rich algae were used in the creation of bio-plastic. In bio-plastic formulation, algae undergo some chemical processing so that their proteins have desired properties. As a side effect of this chemical processing, algae-based bio-products often have unpleasant odors, which must be removed or at least significantly diminished if the products are to be used for commercial purposes. Wang's experiment was carried out to determine the effect of several processing steps on the presence or absence of an unpleasant odor.

Table 1 displays the factors thought to be relevant to bioplastic odor. The researchers designed their experiment investigating the four discrete factors and implemented a design assuming a constant temperature of 25°C. There is currently no known optimal design for such an experiment with both discrete and continuous factors, and this may explain why the researchers at UGA decided to fix the temperature at one arbitrary level. The experiment has a binary response Y denoting whether the odor is successfully removed from the bio-plastic. In this work, we revisit this experiment and consider designs incorporating all four discrete factors and storage temperature as a continuous variable. We model μ , the mean response of Y, using logistic regression and find optimal designs for estimating all parameters in the model. We refer to such studies with both discrete and continuous factors as having "mixed factors." Generalized linear models (GLMs) are widely used to model the mean response of a Bernoulli random variable. Let Y_l be the response corresponding to the *l*th combination of factor levels, $\mathbf{x}_l = (1, \mathbf{x}_{l,1}, \mathbf{x}_{l,2}, ...)$, that may include integration terms among the factors. Without loss of generality, for an experiment with *k* factors, we assume the first term of \mathbf{x}_l corresponds to the intercept, the next *k* terms to the factor settings, and any remaining terms to interactions among the factors. Accordingly, the 2nd through the (k + 1)th terms in \mathbf{x}_l become a support point of the design and the collection of all such points constitutes the experimental design, $\boldsymbol{\psi}$.

In the GLM, the mean response μ_l of Y_l is related to the linear predictor $\eta_l = \mathbf{x}_l^T \boldsymbol{\beta}$ by a monotonic link function $g(\cdot)$ via $g(\mu_l) = \eta_l$, with the logit and probit links being two of the most commonly used when the response is binary. Here, the range of values for each factor is assumed to be known. This implies that we have a known design space where combination levels of the factors can be selected to observe the response.

The exact design problem is to determine the optimal number of support points (*L*), the support points themselves, and the optimal number of replicates, n_1, \ldots, n_L , subject to the constraint $n_1 + \cdots + n_L = N$. The value *N* is the known total number of observations for the study and is predetermined either by the duration or cost of the study. Alternatively, an approximate design optimizes the proportion of the total number of observations at each support point subject to the constraint that they sum to unity (Kiefer 1959). For such designs, we relax the assumption that each proportion $p_l = n_l/N$ is a ratio of two positive integers between 0 and 1 and implement the approximate



 Table 1. Factor types and levels for the bio-plastic odor removal experiment.

		Lev	rels
Туре	Factor	_	+
Discrete Continuous	Algae Scavenger Resin Compatibilizer Temperature	Catfish algae Activated carbon Polyethylene Absent Temperature fre	Solix Microalgae Zeolite Polypropylene Present om 5°C to 35°C

design by taking roughly Np_l observations at each x_l subject to the requirement that they sum to N and each Np_l is an integer.

Let ψ be a design with support points at x_1, \ldots, x_L for which there are n_l replicates at each x_l . A direct calculation shows that the Fisher information matrix is

$$\boldsymbol{I}_{\psi} = \sum_{l=1}^{L} n_l \Upsilon(\eta_l) \boldsymbol{x}_l \boldsymbol{x}_l^T,$$

where $\Upsilon(\eta_l) = \frac{(d\mu_l/d\eta_l)^2}{\mu_l(1-\mu_l)}$. A *D*-optimal design maximizes the log-determinant of the Fisher information matrix and so it is appropriate for estimating all parameters in the model. Because the Fisher information matrix depends on the model parameters, nominal values of the parameters are required before we can implement the design. Such nominal values typically come from the literature or pilot studies. The resulting optimal designs are therefore locally optimal and are often used as building blocks for constructing more complicated designs (Ford, Torsney, and Wu 1992) or as benchmarks for other designs when a single best guess of the parameters is available (Stufken and Yang 2012).

When the design criterion is a concave function of the Fisher information matrix, such as *D*-optimality, we verify the optimality of an approximate design among all designs using an equivalence theorem, see, for example, Kiefer and Wolfowitz (1959) or Pukelsheim (1993). For the logistic model with *q* parameters in the linear predictor, this theorem asserts that the design ψ^* is locally *D*-optimal among all designs if and only if for all *x* in the design space,

$$\frac{\exp\{\boldsymbol{\beta}^T \boldsymbol{x}\}}{(1+\exp\{\boldsymbol{\beta}^T \boldsymbol{x}\})^2} \, \boldsymbol{x}^T \, \boldsymbol{I}_{\psi^*}^{-1} \, \boldsymbol{x} - q \le 0, \tag{1}$$

with equality at each support point of the design ψ^* . The function to the left of the above inequality is sometimes called the sensitivity function.

Often the worth of a design is measured by its efficiency relative to the optimal design, ψ^* . If ψ is a design of interest and ψ^* is a locally *D*-optimal design for a GLM with *q* parameters in the linear predictor, the *D*-efficiency of ψ is

$$\left(\frac{\det(\boldsymbol{I}_{\boldsymbol{\psi}})}{\det(\boldsymbol{I}_{\boldsymbol{\psi}^*})}\right)^{1/q}.$$
(2)

If the ratio is one half, the design ψ requires twice as many replicates as the locally *D*-optimal design to obtain the same information. When the true optimum design is unknown, a lower bound on the *D*-efficiency of ψ is exp $\{-\theta/q\}$, where θ is the maximum positive value of the sensitivity function across the design space (Pazman 1986). Clearly, $\theta = 0$ if and only if ψ is locally *D*-optimal, and the lower bound attains unity. We refer to the quantity in the numerator of (2), det $(I_{\psi})^{1/q}$, as the objective function value for the design and report its value for comparing different designs.

Atkinson and Woods (2015) provided an overview of design issues for generalized linear models. Some theoretical results exist for models with all discrete factors (Yang, Mandal, and Majumdar 2016) or all continuous factors (Yang, Zhang, and Huang 2011). When theoretical results are not available, computational methods are used to find optimal designs. Mandal, Wong, and Yu (2015) provided an overview of algorithms for generating optimal designs, including use of nature-inspired metaheuristic algorithms for finding a large class of optimal designs. Early techniques for generating optimal designs for experiments with outcomes modeled under GLMs include Fedorov-Wynn type algorithms (Fedorov 1972) and multiplicative algorithms (Titterington 1976). These approaches remain popular and often form the basis for more recent techniques such as the cocktail algorithm (Yu 2011). Specific applications of computational methods to solve real design problems for GLMs can be found in Woods et al. (2006), Dror and Steinberg (2006, 2008), Waterhouse et al. (2008), and Woods and van de Ven (2011).

There is little work on constructing efficient designs for experiments with mixed factors; a reason may be that the theory and algorithms for constructing D-optimal designs for GLMs when all factors are continuous or when all factors are discrete do not directly extend to the case when there are mixed factors. As far as we know, there is no efficient algorithm for finding D-optimal designs for such models. Some algorithms, such as quasi-Newton BFGS (Nocedal and Wright 1999), may be used to solve these mixed-factor design problems by optimizing the continuous factor levels and proportions for each fixed combination of discrete factor levels, but such approaches can be computationally inefficient (see Section 3.3). One common approach to these mixed factor problems is to discretize the continuous factors into a few levels and apply algorithms for studies with all discrete factors. However, the generated design is unlikely to be locally D-optimal for the original problem if the discretization is too coarse, as demonstrated in Section 3.2. Haphazard discretization of the continuous factors could also cause separation issues during analysis. This means that valid maximum likelihood estimates of the parameters do not exist because there is a hyperplane in the linear predictors that can perfectly separate the responses into two categories. More complicated and specialized estimation techniques, such as penalized maximum likelihood (Firth 1993; Heinze and Schemper 2002; Woods and van de Ven 2011; Woods, McGree, and Lewis 2017), or a modified logistic regression, such as a hidden logistic regression (Rousseeuw and Christmann 2003), will be required to produce meaningful parameter estimates.

The primary aim of this article is to propose a new method for finding *D*-optimal designs for GLMs with mixed factors and a binary outcome using quantum particle swarm optimization (QPSO). This QPSO is a nature-inspired metaheuristic algorithm based on particle swarm optimization (PSO), which is already widely used in engineering and computer science to tackle complicated optimization problems. A key advantage of working with PSO-type algorithms is that they require only an objective function which can be explicitly written down, and the design space does not have to be discretized. The latter property is particularly useful when we design a study with multiple continuous factors.

This article is organized as follows. Section 2 first provides a brief review of PSO and QPSO before we describe our proposed d-QPSO algorithm. In Section 3, we apply the d-QPSO algorithm to find locally D-optimal designs for several real world problems. Section 4 demonstrates the flexibility of the *d*-QPSO algorithm to find pseudo-Bayesian D-optimal designs for situations in which the nominal values might be unknown. Section 5 summarizes our work with remarks on other possible applications of the *d*-QPSO algorithm. In the supplementary materials, we provide our d-QPSO algorithm code for the odor removal example and show how it may be used to investigate robustness properties of the *D*-optimal designs to violation of the model assumptions for the motivating example. We report computational time and accuracy of the *d*-QPSO-generated designs via simulations and also demonstrate that the proposed d-QPSO algorithm can be used to find exact optimal designs.

2. Swarm Optimization

We begin by briefly reviewing PSO and QPSO. We then describe how we modify QPSO to *d*-QPSO for finding *D*-optimal designs for models with mixed factors and a binary response.

2.1 Particle Swarm Type Algorithms

PSO is a metaheuristic optimization algorithm introduced by Kennedy and Eberhart (1995). It is a nature-inspired algorithm that mimics the behavior of a flock of birds as they search an area for food. Each member of the flock or swarm, known as a particle, represents a candidate solution to the problem of interest with a corresponding fitness, and the location of the food represents the optimum solution. Each particle has its own perception of where the food is located, based on its own experience. This position is known as the personal best position (pbest). Each particle is also aware of the overall best location that the flock has found, a position called the global best position (gbest). At each iteration, every particle moves in the direction of both its pbest position and the gbest position.

Since its inception, many variants of PSO have been developed, often to adapt PSO to a specific class of problems. For example, in public health research, Fu et al. (2009) used PSO to identify optimal screening nodes for spread of the SARS disease in Singapore; other applications are voluminously documented in the engineering literature, such as in the IEEE Transactions. Given its success in other application areas, PSO has also been modified to find optimal designs. Qiu et al. (2014) appeared to be the first to use the standard PSO to find a variety of optimal designs for biomedical problems, including optimal designs for estimating parameters in compartmental models and tumor growth models. It has also been used to find optimal designs under a nondifferentiable optimality criterion (Chen et al. 2015a), optimal designs for a variety of mixture models (Wong et al. 2015), optimal latin hypercube designs (Chen et al. 2013), and most recently, minimax projection designs (Mak and Joseph 2017). While the standard PSO is fast and effective for finding optimal designs for a few factors, it may not work very efficiently for complicated design problems, such as the case when we have moderate to large number of mixed factors in the regression model. This leads us to explore QPSO as a general optimization algorithm before modifying it to d-QPSO to specifically solve difficult design problems.

QPSO was developed after the trajectory analysis by Clerc and Kennedy (2002) where they showed that the swarm converges if each particle converges to the local attractor, which is defined as a point between the pbest and gbest positions in the standard PSO algorithm. QPSO was first introduced by Sun, Feng, and Xu (2004a), with the central idea that each particle can appear anywhere in the search space at any time, but has a higher probability to appear near its current position. This probabilistic scheme is unique to QPSO and justifies the use of the term "Quantum" in its name. Unlike the standard PSO, QPSO has no velocity term in its defining equations. Each particle's stochastic movement is accomplished by drawing positions from an exponential distribution with parameters determined by the distance between the particle and the best known positions. This probabilistic draw incorporates the local attractor for each particle and a position known as the "mainstream thought" or mbest, which is the average of all pbest positions at the current iteration (Sun, Xu, and Feng 2004b). By updating particle positions using both the local attractor and mbest (see Figure 1), QPSO is able to draw particles toward optimal positions without throwing away information from the particles with poor fitness values.

To our knowledge, this is the first work to apply QPSO to design experiments. The algorithm cannot simply be used off-the-shelf to find optimal designs because it assumes an unbounded search space and generally does not optimize discrete and continuous factors simultaneously. Thus, we develop a variant of QPSO for design of experiments and call it *d*-QPSO for short, where d stands for design. In particular, we modify the algorithm's behavior and use an elitist breeding mutation technique to maximize its performance for finding our sought-after designs. The overall idea of the d-QPSO algorithm is to generate multiple quantum-behaved particle swarms, where each swarm works to find the global best position. Information within swarms is shared just like in a typical QPSO algorithm, but with special attention paid to the nature of the covariates, the model structure, and the type of design space. Information is also pooled across swarms via an elitist breeding mutator without sacrificing the possibly distinct solutions obtained from different swarms.

2.2 d-QPSO: Algorithm Overview

Our proposed algorithm is a multi-swarm QPSO with elitist breeding that proceeds as follows. Suppose the design problem for our binary response experiment has k factors, and, to fix ideas, assume that all discrete factors have two levels. We first randomly generate s swarms each with w particles, where each particle is a design with L support points, and s, w, and L are user-selected positive integers. We refer to the collection of all swarms as a habitat. The components of each particle are the L support point settings and the proportion of observations at each support point. Thus each particle, ψ_i , $i = 1, \ldots w$ has L(k + 1) elements and we search for an optimal design among them. Within each swarm, $\boldsymbol{\psi}_i$, i = 1, ..., w is generated randomly at the start of the search. We denote the best set of factor settings and proportion allocations found by particle *i* at iteration *t* by $\boldsymbol{\psi}_i^{(\text{pbest},t)} = (\boldsymbol{\psi}_{i,1}^{(\text{pbest},t)}, ..., \boldsymbol{\psi}_{i,L(k+1)}^{(\text{pbest},t)})^T$, and the best set of factor settings and proportion allocations found by the entire swarm by $\boldsymbol{\psi}^{(\text{gbest},t)} = (\boldsymbol{\psi}_1^{(\text{gbest},t)}, ..., \boldsymbol{\psi}_{L(k+1)}^{(\text{gbest},t)})^T$. At t = 0, the pbest positions are the initial positions of each particle and the gbest position is the best of all pbest positions, where "best" refers to the position corresponding to the design with the largest value of the objective function. Each iteration has two update steps, one at the swarm-level and the other at the habitat-level (Figure 1).

Below we enumerate the steps in the *d*-QPSO algorithm for updating each swarm and suppress the swarm indicator for notational simplicity. The updates must be performed differently for continuous factor settings, discrete factor settings, and the proportion allocated to each support point. We collect the components of each ψ_i into three sets, A_c , A_d , and A_p which correspond to continuous factors, discrete factors, and proportions, respectively. Clearly, $A_c \cup A_d \cup A_p = \{1, 2, ..., L(k+1)\}$, and updating each set updates the entire vector.

1. Update the local attractors. For each particle *i*, *i* = 1, ..., w, the local attractor $\mathbf{a}_i^{(t)} = (a_{i,1}^{(t)}, ..., a_{i,L(k+1)}^{(t)})^T$ is the central point around which the particle will appear at iteration *t*. For a given swarm, the local attractor for particle *i* at component *m*, *m* = 1, ..., L(k+1) is calculated as

$$a_{i,m}^{(t)} = \begin{cases} \phi_{i,m}^{(t)} \times \psi_{i,m}^{(\text{pbest},t-1)} & \\ +(1 - \phi_{i,m}^{(t)}) \times \psi_m^{(\text{gbest},t-1)} & \text{if } m \in A_c \cup A_p, \\ \psi_m^{(\text{pbest},t-1)} & \text{if } m \in A_d \text{ and } r_{i,3} \le 0.5, \\ \psi_m^{(\text{gbest},t-1)} & \text{if } m \in A_d \text{ and } r_{i,3} > 0.5, \end{cases}$$

where $\phi_{i,m} = \frac{r_{i,1}}{r_{i,1}+r_{i,2}}$ and $r_{i,1}, r_{i,2}, r_{i,3}$ are independent draws from U(0, 1), the uniform distribution over the interval (0, 1).

2. Update particle positions. Each particle is drawn to both its local attractor, $\mathbf{a}_i^{(t)}$ and to $\boldsymbol{\psi}^{(\text{mbest},t)} =$

 $(\frac{1}{w}\sum_{i=1}^{w}\psi_{i,1}^{(\text{pbest},t-1)},\ldots,\frac{1}{w}\sum_{i=1}^{w}\psi_{i,L(k+1)}^{(\text{pbest},t-1)})^{T}$, a position referred to as the "mbest." For discrete factors, that is, $m \in A_d$, $\psi_m^{(\text{mbest},t)}$ is rounded to 1 or -1. The position update step is

$$\psi_{i,m}^{(t)} = \begin{cases} a_{i,m}^{(t)} + \alpha \big| \psi_{i,m}^{(t-1)} \\ -\psi_m^{(mbest,t)} \big| \log(r_{i,m,1}) & \text{if } m \in A_c \cup A_p \text{ and } r_{i,m,2} \ge 0.5 \text{ ,} \\ a_{i,m}^{(t)} - \alpha \big| \psi_{i,m}^{(t-1)} \\ -\psi_m^{(mbest,t)} \big| \log(r_{i,m,1}) & \text{if } m \in A_c \cup A_p \text{ and } r_{i,m,2} < 0.5 \text{ ,} \\ -\psi_{i,m}^{(t-1)} & \text{if } m \in A_d \text{ with probability } \tau_l, \\ \psi_{i,m}^{(t-1)} & \text{if } m \in A_d \text{ with probability } 1 - \tau_l, \end{cases}$$

where $r_{i,m,1}$, $r_{i,m,2}$ are independent U(0, 1) draws. Here α is known as the contraction-expansion coefficient and is decreased linearly from 1.4 to 0.4 as the algorithm runs (Sun et al. 2012). For the discrete factor update, τ_l is the probability of changing factor setting for factors belonging to x_l and is calculated using Hamming distance as described in Xi et al. (2016). Under this updating scheme, a vector of discrete factor settings is created by combining $\boldsymbol{\psi}^{(\text{mbest},t)}$ and $\boldsymbol{\psi}^{(\text{gbest},t-1)}$ using a crossover. In a crossover, two position vectors are split at a random location and the left partition of one is combined with the right partition of the other to form a new position vector. We use the vector with left elements taken from $\boldsymbol{\psi}^{(\text{mbest},t)}$ and right elements taken from $\boldsymbol{\psi}^{(\text{gbest},t-1)}$ and refer to this vector as $\boldsymbol{\psi}^{(\text{cbest},t)}$. The larger the Hamming distance between $\psi^{(\text{cbest},t)}$ and the particle's current position, the more likely each element of $\psi_i^{(t-1)}$ is to "flip" factor settings. For discrete factors, $m \in A_d$, τ_l is calculated based on the x_l to which *m* belongs. That is, for support point x_l , the probability of flipping, τ_l , will be the same for all discrete factor settings of that support point. The value of τ_l is

 $\tau_l = \min\left\{\frac{\alpha \times b_l(-\log(r))}{k_d}, 1\right\},\,$



Figure 1. Steps in the QPSO update for generating locally D-optimal approximate designs. The swarm update is applied to each swarm individually, and the habitat update is performed on all swarms.

where k_d is the number of discrete factors in the model, r is a random draw from U(0, 1), and b_l counts the number of elements in the current candidate $\boldsymbol{\psi}_i^{(t)}$ which are different from the corresponding elements of $\boldsymbol{\psi}^{(\text{cbest},t)}$, restricted to the discrete factors only, for \boldsymbol{x}_l . The probability of changing a factor setting, τ_l , increases as b_l increases. After the position update, any proportions that fall below 0 are set to 0. The proportions are then normalized to sum to one. If any value of a continuous factor falls outside its range, we assign this value to the nearest boundary on the design space.

- 3. Update fitness. For each particle *i*, i = 1, ..., w in a given swarm, we calculate the value of the objective function. If a locally *D*-optimal design is sought, the fitness function is log det(I_{ψ_i}), where the logarithm is used for numerical stability and to ensure the design criterion is concave for the application of the equivalence theorem. The fitness function used to find pseudo-Bayesian designs is the approximation to the expected log-determinant of the Fisher information matrix developed in Gotwalt, Jones, and Steinberg (2009).
- 5. *Elitist breeding*. To encourage exploration of the search space, we include an elitist breeding mutator similar to the one described in Yang, Wu, and Min (2015). At the end of each iteration, each particle undergoes elitist breeding with a small probability (we use probability 0.1). A new position vector is created $\boldsymbol{\psi}_i^{(\text{pbest}^*)} = \boldsymbol{\psi}_i^{(\text{pbest},t)}$ and updated by replacing a randomly selected element, *m* with a value taken from another randomly selected element, *m** of a random particle's pbest. Here both *m* and *m** belong to $\{1, 2, \ldots, L(k+1)\}$ and $\boldsymbol{\psi}_{i,m}^{(\text{pbest}^*)}$ is updated as

$$\psi_{i,m}^{(\text{pbest}^*)} = \psi_{i^*,m^*}^{(\text{pbest},t)},$$

where $i^* \in \{1, \ldots, w\}$, and the position $\psi_{i^*, m^*}^{(\text{pbest}, t)}$ is normalized to be in the range of component *m*. Figure 1 includes a schematic diagram of the elitist breeding mutator.

Following the swarm update, the habitat update keeps track of the best position from each swarm and the overall best position, which we denote by $\psi^{(\text{hbest},t)}$, where hbest is short for "habitat best." We share information between swarms using another elitist breeding mutator. This acts to prevent the swarms from becoming stuck in local extrema. This elitist breeding update is similar to the swarm update, except that instead of selecting particles to breed from within the same swarm, particles are allowed to breed across swarms. To our knowledge, this is the first QPSO algorithm to share information across several swarms using such a technique. At each iteration, we determine whether the hbest position has improved over all previous iterations. If we are searching for a locally *D*-optimal approximate

design, we check the local *D*-optimality of this generated design by examining if its sensitivity function satisfies the equivalence theorem. This can be done either by a grid search or another QPSO algorithm to maximize the sensitivity function and determine its minimum *D*-efficiency lower bound. The algorithm terminates if the desired *D*-efficiency lower bound is obtained or the maximum number of iterations has been reached. At termination, the design constructed using $\psi^{(\text{hbest},t)}$ is returned as the *d*-QPSO algorithm-generated design.

Tuning the *d*-QPSO algorithm involves selecting the number of particles per swarm, the total number of swarms, the maximum number of iterations, and the maximum number of support points allowable in the design. For locally D-optimal approximate designs, the user must also supply a lower bound for the D-efficiency of the generated design; the algorithm terminates if the lower bound is met. Our general experience is that a relatively small number of particles (around w = 30) works well, allowing each particle to have a strong influence on the mean best position. As a rule of thumb, we suggest using k, swarms, where k is the number of factors in the model. We also suggest increasing the number of swarms when there are more factor levels in the study. The maximum number of iterations is decided by the user and this number should increase with the number of factors. In general, we find that a few thousand iterations work quite well. Finally, for approximate designs the maximum number of support points should be chosen based on the number of factors in the model. For small number of factors, it is appropriate to allow $L = 2^k$ support points; the *d*-QPSO algorithm will generally be able to find a design supported on fewer points. Our experience is that for many experiments with q parameters in the model, having particles each with 2q to 3q points in the algorithm seems to work well. Of course, if these numbers fail to provide an adequate design, the number of support points can always be increased.

We have run extensive simulations to assess the effectiveness of the *d*-QPSO algorithm. These results, along with several other sets of results obtained from the *d*-QPSO algorithm, are provided in the supplementary materials. In Section S1, we provide CPU time and accuracy simulations for designing mixed factor experiments using the *d*-QPSO algorithm. Section S2 uses the *d*-QPSO algorithm to find minimally supported designs. All computations in this article were carried out using a 2012 Macbook Pro 2.6GHz Intel Core i7 with 16G RAM on 64bit OSX EI Capitan. Our code is written in C++ and called from R via the RCPP package (Eddelbuettel and François 2011).

3. Applications: Locally D-Optimal Designs

We now demonstrate that the *d*-QPSO algorithm can find locally *D*-optimal designs for GLMs with mixed factors and a binary response of increasing complexity. We start by revisiting the motivating odor removal example and find a more realistic locally *D*-optimal approximate design for the study. We then apply *d*-QPSO and obtain locally *D*-optimal approximate designs for an electrostatic discharge experiment and a 10-factor car refueling experiment. In all examples, we use the logit link function in the GLM, but other link functions can be used as well (see the supplementary materials Section S3.1).

Table 2. The	e d-QPSO	algorithm-generated	locally	D-optimal	approximate	design	for	the	odor	removal	experiment	with	nominal	values	$\boldsymbol{\beta} = (-$	-1, 2, 0.5,	-1,
-0.25, 0.13)	т.																

Support point	Algae	Scav.	Resin	Comp.	Temp.	<i>p_i</i> (%)	Support point	Algae	Scav.	Resin	Comp.	Temp.	p _i (%)
1	-1	-1	-1	-1	9.040	3.70	8	-1	1	1	-1	16.894	2.20
2	-1	-1	-1	-1	25.788	4.30	9	-1	1	1	-1	33.366	8.80
3	-1	-1	-1	1	29.710	10.17	10	-1	1	1	1	35.000	6.10
4	-1	-1	1	-1	35.000	4.73	11	1	-1	-1	1	5.000	5.11
5	-1	-1	1	1	29.579	11.59	12	1	-1	1	-1	5.000	10.75
6	-1	1	-1	-1	5.000	9.80	13	1	-1	1	1	5.000	5.23
7	-1	1	-1	1	5.206	7.86	14	1	1	1	1	5.000	9.71

3.1 Odor Removal Experiment

In the motivating odor removal study conducted by Wang et al. (2016), the experimenters fixed the temperature variable and used a 2⁴⁻¹ regular fractional factorial design with equal number of replicates. The *d*-QPSO algorithm allows us to construct a design that includes storage temperature as a continuous factor for the model $Y \sim \text{Bern}(\mu)$ with $\text{logit}(\mu) = \beta_0 + \beta_1 \text{Algae} + \beta_2 \text{Scavenger} + \beta_3 \text{Resin} + \beta_4 \text{Compatibilizer} + \beta_5 \text{Temperature}$. Using information from their study, we take nominal values $\boldsymbol{\beta} = (-1, 2, 0.5, -1, -0.25, 0.13)^T$ and apply the *d*-QPSO algorithm to construct a locally *D*-optimal approximate design. The tuning parameters we used were 4 swarms, 30 particles in each swarm, and we initialized our search among designs with up to 20 support points. The termination rule was either a maximum of 6000 iterations or when the generated design attained a *D*-efficiency lower bound of 99%.

Table 2 displays the *d*-QPSO algorithm-generated locally *D*-optimal approximate design with 14 support points, where only some of the points have values at the extreme ends of the temperature range. The objective function value for this design is $0.0019^{1/6} = 0.3519$. The sensitivity plot of the design in Figure 2 confirms its *D*-optimality. It is instructive to compare this design to three locally *D*-optimal approximate designs obtained by breaking up the continuous temperature variable into 2, 3, and 4 uniformly spaced points across the design space. Designs using the discretized temperature variable were obtained using the lift-one algorithm (Yang, Mandal, and Majumdar 2016). The *D*-efficiencies of these three designs relative to the *d*-QPSO algorithm-generated design were 0.9737, 0.9907, and 0.9965,

but the minimum allocations required at a support point were 0.46%, 1.54%, and 0.35% for the 2, 3, and 4 uniformly spaced point designs, respectively. This means that although theoretically these designs can be highly efficient, unless the total number of observations in the study is large, these lift-one algorithm-generated designs cannot be implemented in practice. In particular, the four-point lift-one algorithm-generated design requires 0.0035N observations at a support point versus 0.022N for the *d*-QPSO algorithm generated design.

3.2 Electrostatic Discharge Experiment

Whitman et al. (2006) described a similar experiment with continuous and discrete factors. The experimenters were interested in finding factors that influence the failure of semiconductors when exposed to electrostatic discharge (ESD). The response was whether or not a certain part of the semiconductor failed, and the model was logistic regression with five factors. The first two factors, Lot A and Lot B, describe the type of wafer used. The third factor was ESD handling: whether or not proper ESD precautions were taken. ESD testing requires a part to be "zapped" with a pulse at either a positive or a negative polarity and then to be zapped again by a pulse with the opposite polarity. A lack of standardization of which pulse order should be used resulted in the experimenters using pulse order as the fourth factor. The final factor, voltage applied to the chip, was continuous and ranged from 25 to 45. The experimental units were single chip TDMA power amplifiers, chosen for their lack of ESD protection circuitry. Without protection circuitry, the chips were expected to be sensitive to changes in experimental factors. Taking Y to



Temperature

Figure 2. Sensitivity plot of the *d*-QPSO algorithm-generated locally *D*-optimal approximate design for each fixed combination of discrete factors in the odor removal experiment.

Table 3. The *d*-QPSO algorithm-generated locally *D*-optimal approximate design for the ESD experiment with nominal values $\boldsymbol{\beta} = (-7.5, 1.50, -0.2, -0.15, 0.25, 0.35, 0.4)^T$.

Support point	A	В	ESD	Pulse	Volt.	p _i (%)	Support point	A	В	ESD	Pulse	Volt.	p _i (%)
1	-1	-1	-1	-1	28.04	1.80	8	-1	1	-1	1	25.00	10.00
2	-1	-1	-1	-1	25.00	7.46	9	-1	1	1	-1	25.00	3.80
3	-1	-1	-1	1	25.00	2.49	10	-1	1	1	-1	32.93	13.43
4	-1	-1	-1	1	27.85	7.74	11	-1	1	1	1	25.00	9.20
5	-1	-1	1	-1	25.00	11.65	12	1	-1	1	-1	25.00	1.23
6	-1	-1	1	1	25.00	8.58	13	1	1	1	-1	25.00	13.40
7	-1	1	-1	-1	25.00	9.20							

be 1 if the chip fails and 0 otherwise, we have $Y \sim \text{Bern}(\mu)$, and the model of interest is $\text{logit}(\mu) = \beta_0 + \beta_1 \text{Lot A} + \beta_2 \text{Lot B} + \beta_3 \text{ESD} + \beta_4 \text{Pulse} + \beta_5 \text{Voltage} + \beta_{34} (\text{ESD} \times \text{Pulse}).$

The experimenters discretized the voltage variable into five levels: 25, 30, 35, 40, and 45 V. The full factorial design with 80 separate settings was implemented to test all possible combinations of factor settings. They did not provide a rationale for treating voltage as a discrete variable, but we note that such techniques are common when constructing exact response surface designs with continuous factors.

We redesign this experiment by treating voltage as a continuous variable that ranges from 25 to 45 instead of fixing it to 5 levels. We use the set of nominal values $\boldsymbol{\beta} =$ $(-7.5, 1.50, -0.2, -0.15, 0.25, 0.35, 0.4)^T$, motivated by the parameter estimates reported in the article. The tuning parameters we used were 5 swarms, 30 particles per swarm, and we initialized our search among designs with up to 18 support points. The termination rule was either a maximum of 4000 iterations or a generated design that attains a *D*-efficiency lower bound of 99%. Table 3 lists the *d*-QPSO algorithm-generated design with 13 points and its sensitivity function in Figure 3 confirms its local *D*-optimality.

The *d*-QPSO algorithm-generated locally *D*-optimal approximate design has four unique voltage settings. The original experimental design has five unique voltage settings, and these were very different than the ones identified by the *d*-QPSO algorithm, which has an objective function value of $(1.2639 \times 10^{-5})^{1/7} = 0.1997$. The *D*-efficiency of the original 80 point design relative to the design identified by the *d*-QPSO algorithm is 32.85%, indicating that the latter design is approximately three times as *D*-efficient as the implemented design.

3.3 Car Refueling Experiment

We now apply the *d*-QPSO algorithm to solve a highdimensional design problem. Grimshaw et al. (2001) described an experiment to test a vision-based car refueling system. Here, the investigators were interested in finding whether a computer-controlled nozzle was able to insert itself into the gas pipe correctly or not. Table 4 lists the four discrete factors and the six factors that are naturally continuous: reflective ring thickness, lighting angle, gas-cap angle (*Z*axis), gas-cap angle (*Y*-axis skew), car distance, and threshold step value. For illustrative purposes, we take the set of nominal values for the model parameters to be: $\boldsymbol{\beta} =$ (3, 0.5, 0.75, 1.25, 0.8, 0.5, 0.8, -0.4, -1.00, 2.65, 0.65)^T, where the order of the factors is the order given in Table 4. Other sets of nominal values could also be used.

We use the *d*-QPSO algorithm to search for a locally *D*-optimal approximate design for the main effects logistic model. The tuning parameters we used were 10 swarms, 30 particles per swarm, and we initialized our search among designs with up to 12 support points. The termination rule was a maximum of 7000 iterations or a generated design attaining a *D*-efficiency lower bound of 99%, which was determined using a second QPSO algorithm to find the maximum of the sensitivity function. This second search was used because the design has six continuous factors, and thus performing a grid search to find the maximum of the sensitivity function is difficult. While this search did not find any values to indicate that our design was not locally *D*-optimal via the equivalence theorem, we note that there is a possibility that we did not find the maximum of the sensitivity function that determines the minimum *D*-efficiency



Voltage

Figure 3. Sensitivity plot of the d-QPSO algorithm-generated locally D-optimal approximate design for each fixed combination of discrete factors in the ESD experiment.

Table 4. Factor types and levels for the car refueling experiment.

			Levels
Туре	Factor	Low	High
Discrete	Ring type Lighting Sharpen Smooth	White paper Room lighting No No	Reflective 2 flood lights and room lights Yes Yes
Continuous	Lighting angle Gas-cap angle (Z-axis) Gas-cap angle (Y-axis skew) Car distance Reflective ring thickness Threshold step value	from 50 from 30 from 0 from from	degrees to 90 degrees degrees to 55 degrees degrees to 10 degrees om 18 in. to 48 in. 0.125 in. to 0.425 in. from 5 to 15

Table 5. The *d*-QPSO algorithm-generated approximate design for the car refueling experiment with nominal values $\beta = (3, 0.5, 0.75, 1.25, 0.8, 0.5, 0.8, -0.4, -1.00, 2.65, 0.65)^{T}$.

Ring type	Lighting	Sharpen	Smooth	Lighting angle	Z-axis	Y-axis skew	Car dist.	Ring thick.	Threshold stepsize	p _i (%)
-1	- 1	-1	-1	50.00	30.00	4.20	48.00	0.125	5.00	9.1
-1	-1	-1	-1	50.00	30.00	10.00	48.00	0.125	8.57	9.1
—1	-1	-1	-1	50.00	30.00	10.00	45.68	0.125	5.00	9.1
—1	-1	-1	-1	54.64	30.00	10.00	48.00	0.125	5.00	9.1
—1	-1	-1	-1	50.00	32.90	10.00	48.00	0.125	5.00	9.1
—1	-1	-1	-1	50.00	30.00	10.00	48.00	0.125	5.00	8.1
—1	-1	-1	-1	50.00	30.00	10.00	48.00	0.425	5.00	7.7
—1	-1	-1	1	50.00	30.00	10.00	48.00	12.5	5.00	9.1
—1	-1	1	-1	50.00	30.00	10.00	48.00	12.5	5.00	9.1
—1	1	-1	-1	50.00	30.00	10.00	48.00	12.5	5.00	9.1
1	-1	-1	-1	50.00	30.00	10.00	48.00	12.5	5.00	7.5
1	-1	-1	-1	50.00	30.00	10.00	48.00	0.425	5.00	4.0

of the design generated. Table 5 displays the *d*-QPSO algorithmgenerated design with an objective function value of $(2.5181 \times 10^{-16})^{1/11} = 0.0382$.

To further support our claim that this is design highly *D*-efficient, and is indeed locally *D*-optimal, we run the *d*-QPSO algorithm several more times, holding the number of swarms fixed at 10 and changing the number of particles, the limit on the number of possible support points, and the maximum number of iterations. Table 6 provides a summary of the tuning parameters used and the results obtained. In all cases, we obtain a design similar to the one presented in Table 5. Even the worst design, obtained using only 10 particles per swarm, still had a *D*-efficiency of 98% relative to the *d*-QPSO algorithm-generated locally *D*-optimal approximate design in Table 5.

We also attempted to solve this 10-factor problem by discretizing the design space and using current algorithms, such

 Table 6. The *d*-QPSO algorithm tuning parameters and results for constructing locally *D*-optimal approximate designs for the car refueling experiment.

Max support points	Particles per swarm	Max iterations	log det	Support points	CPU time (Seconds)
12	35	5000	- 35.91	12	87.58
12	35	7000	- 36.08	11	49.38
14	35	5000	- 36.04	11	47.51
14	35	7000	- 36.02	11	244.70
14	20	7000	- 36.01	12	140.80
16	35	5000	- 35.93	12	70.44
16	10	7000	- 36.15	12	67.12
16	20	7000	- 35.98	12	85.60
16	35	7000	- 35.92	12	211.34
16	75	7000	- 35.93	12	424.52

as the lift-one and Fedorov–Wynn type of algorithms; however, the number of candidate points became too large due to the number of factors, and the algorithms were unable to run successfully. We also applied quasi-Newton BFGS to tackle this problem via the optim package in R. This algorithm cannot search over the discrete factor settings, so we resorted to fixing the discrete factor combinations and applying quasi-Newton BFGS to find the continuous factor settings along with the proportions allocated to each support point. The algorithm took several hours to run, significantly longer than the *d*-QPSO algorithm. The resulting design had 16 support points and its *D*-efficiency was only 70% relative to the *d*-QPSO algorithm-generated locally *D*-optimal approximate design in Table 5.

4. Applications: Pseudo-Bayesian Designs

The previous section assumes that we have a given set of nominal values for the model parameters to find an optimal design. When no such reliable nominal values are available, Bayesian designs provide an attractive, robust solution to the design problem. These designs require priors for the parameter values and proceed by maximizing the expectation of the log-determinant of the Fisher information matrix by choice of a design. In general, Bayesian designs come at a high computational cost, as numerical integration must be performed every time the design changes. To circumvent these numerical difficulties, several pseudo-Bayesian methods have been proposed. Woods et al. (2006) identified exact compromise designs that are robust to misspecification of both the parameter values and link function.

Table 7. The *d*-QPSO algorithm-generated pseudo-Bayesian approximate design for the odor removal experiment using the independent uniform priors for all parameters specified in Section 4.1.

Support point	Algae	Scav.	Resin	Comp.	Temp.	p _i (%)	Support point	Algae	Scav.	Resin	Comp.	Temp.	p _i (%)
1	-1	-1	-1	-1	5.000	8.50	8	-1	1	1	1	14.674	5.00
2	-1	-1	-1	-1	34.911	3.87	9	-1	1	1	1	34.980	9.49
3	-1	-1	-1	1	28.818	6.60	10	1	-1	-1	1	5.000	9.80
4	-1	-1	1	-1	31.261	10.10	11	1	-1	1	— 1	5.000	4.70
5	-1	1	-1	<u> </u>	5.000	7.60	12	1	-1	1	1	5.000	9.92
6	-1	1	-1	<u> </u>	21.550	3.15	13	1	1	1	<u> </u>	5.000	10.79
7	— 1	1	-1	1	5.000	8.18	14	1	1	1	1	5.000	2.30

Table 8. The *d*-QPSO algorithm-generated pseudo-Bayesian exact design with 16 support points for the crystallography experiment with independent prior specification $\beta_0 \sim U(-3, 3), \beta_1 \sim U(4, 10), \beta_2 \sim U(5, 11), \beta_3 \sim U(-6, 0), \text{ and } \beta_4 \sim U(-2.5, 3.5).$

Support point	Agitation rate	Composition volume	Temperature	Evaporation rate	Support point	Agitation rate	Composition volume	Temperature	Evaporation rate
1	- 1.000	0.378	- 1.000	- 1.000	9	0.594	- 1.000	- 1.000	- 1.000
2	- 1.000	0.791	-1.000	- 1.000	10	1.000	- 0.990	- 1.000	- 1.000
3	- 1.000	0.212	- 1.000	1.000	11	0.446	- 1.000	- 1.000	1.000
4	- 1.000	0.708	- 1.000	1.000	12	0.940	- 1.000	- 1.000	1.000
5	- 0.969	1.000	1.000	- 1.000	13	1.000	- 0.652	1.000	- 1.000
6	- 0.487	1.000	1.000	- 1.000	14	1.000	- 0.223	1.000	- 1.000
7	- 1.000	1.000	1.000	1.000	15	1.000	- 0.817	1.000	1.000
8	- 0.522	1.000	1.000	1.000	16	1.000	— 0.397	1.000	1.000

Gotwalt, Jones, and Steinberg (2009) used a quadrature scheme to approximate the value of the expected Fisher information matrix. Overstall and Woods (2017) developed an approximate coordinate exchange algorithm for identifying Bayesian designs through the use of a Gaussian process emulator. We next demonstrate the *d*-QPSO algorithm's ability to find pseudo-Bayesian designs using the odor removal experiment and the 16 run crystallography experiment described by Woods et al. (2006).

4.1 Odor Removal Experiment

We first return to the odor removal experiment. We use the d-QPSO algorithm to obtain a robust pseudo-Bayesian design by assuming independent uniform priors for each parameter with each prior centered at the nominal value and having a width twice the magnitude of the supposed nominal value used in Section 3.1. For example, if the nominal value of a parameter was 1, we use the uniform prior over (0, 2). These independent priors for β_0, \ldots, β_5 are: U(-2, 0) for β_0 and $\beta_3, U(0, 4)$ for β_1 , U(0, 1) for β_2 , $U(-\frac{1}{2}, 0)$ for β_4 , and U(0, 0.26) for β_5 . The tuning parameters we used were 2 swarms, 30 particles per swarm, and we initialized our search among designs with up to 16 support points. The termination rule was 10,000 iterations. Table 7 displays the *d*-QPSO algorithm-generated design. We compare robustness properties of this design with the *d*-QPSO algorithm-generated locally D-optimal approximate design in Section S3.2 of the supplementary materials.

4.2 Crystallography Experiment

Next we consider obtaining a pseudo-Bayesian design for the crystallography experiment described in Woods et al. (2006). Following Gotwalt, Jones, and Steinberg (2009), we use the prior

labeled β_3 in Table 1 of Woods et al. (2006). This prior specification is $\beta_0 \sim U(-3, 3)$, $\beta_1 \sim U(4, 10)$, $\beta_2 \sim U(5, 11)$, $\beta_3 \sim$ U(-6, 0), and $\beta_4 \sim U(-2.5, 3.5)$. The experiment has four continuous factors corresponding to the agitation rate, composition volume, temperature, and evaporation rate. The *d*-QPSO algorithm tuning parameters we used were 5 swarms, 30 particles per swarm, and we initialized our search for the optimal exact design among all optimal exact designs with up to 16 support points. The termination rule was 7000 iterations. Table 8 displays the *d*-QPSO algorithm-generated design for this optimal exact design problem. For comparison, we generate one million randomly sampled parameter vectors from the prior and calculate the objective function value of the d-QPSO algorithmgenerated design, the design in Woods et al. (2006), the design presented by Gotwalt, Jones, and Steinberg (2009), the design obtained using the method of Gotwalt, Jones, and Steinberg (2009) reported in Overstall and Woods (2017) (referred to as OW Gotwalt), and the design obtained using the approximate coordinate exchange (ACE) algorithm in the acebayes R package from Overstall and Woods (2017). Table 9 compares these methods using the average objective function values and the median D-efficiencies of the designs relative to the d-QPSO algorithmgenerated design. The table shows that the *d*-QPSO algorithm is competitive with the other methods.

Table 9. Mean objective function value and median *D*-efficiency relative to the *d*-QPSO algorithm-generated pseudo-Bayesian exact design (RE) for each of the robust designs considered for the crystallography experiment over one million randomly sampled parameter vectors.

Design	Mean $ X^TWX ^{1/5}$	Median RE
<i>d</i> -QPSO	0.5734	1.0000
Gotwalt, Jones, and Steinberg (2009)	0.4643	0.8194
OW Gotwalt	0.5687	0.9896
Woods et al. (2006)	0.5394	0.9373
ACE	0.5685	0.9933

5. Summary and Conclusions

In this article, we proposed a novel and flexible d-QPSO algorithm to find several types of D-optimal designs for GLMs with mixed factors and a binary response. We demonstrated that the d-QPSO algorithm could find more D-efficient designs than those obtained by treating continuous factors as discrete. We applied the d-QPSO algorithm to find a locally D-optimal design for an experiment with 10 factors, and we also showed it can be used to find robust pseudo-Bayesian designs when there is uncertainty in the parameter values.

In conclusion, we view the use of metaheuristic optimization algorithms as an effective option for finding solutions to complicated design problems. We believe that there is definite potential for further use of nature-inspired metaheuristic algorithms to find different kinds of optimal designs and help us better understand properties of optimal designs. For example, Section S3.1 of the supplementary materials provides a brief study on sensitivity of the locally D-optimal designs to misspecification of the link function. Applying the d-QPSO algorithm to find optimal exact designs for correlated responses or minimum bias designs is potentially interesting. For such problems, there are no equivalence theorems to resort to for confirming optimality of the generated design because the design criteria are no longer concave. The only way to assess the optimality of the generated design by the *d*-QPSO algorithm or other algorithms is by developing theoretical results, which are usually only feasible for relatively simple models. A plausible strategy in this situation is to show the algorithm generates the same optimal designs that are already worked out analytically for simple cases, and then use the algorithm to find optimal designs for more complicated cases where theoretical designs are no longer available. For example, Chen et al. (2015b) used PSO to generate locally Doptimal exact designs for the Michaelis-Menten model with correlated errors and confirm their numerical results with the theoretical optimal designs available from Dette and Kunert (2014) when only a couple of time points are allowed for taking measurements. Chen et al. (2015b) then used PSO to generate optimal designs for a longitudinal study with more time points than those considered in Dette and Kunert (2014), where theoretical results are not available.

Supplementary Materials

- **S1.** *d***-QPSO Computational Timing and Accuracy:** Simulations tracking the speed and accuracy of the *d*-QPSO algorithm for finding locally *D*-optimal approximate designs for models with up to six mixed factors and a binary response.
- **S2. Minimally Supported Designs:** Simulations using the *d*-QPSO algorithm to construct locally *D*-optimal approximate designs for models with one discrete and one continuous factor and a binary response. We investigate conditions under which the *d*-QPSO algorithm can find minimally supported designs.
- **S3.** Sensitivity Study: (i) An illustrative study of the robustness of two factor experiments with binary response to mis-specification of the link function and (ii) a study of

the robustness of the optimal design for the odor removal experiment to parameter mis-specification.

- **S4.** Locally *D*-optimal Exact Designs: Selected *D*-optimal exact designs for the odor removal experiment for various total sample sizes *N*.
- **S5.** Optimal Designs on an Irregular Design Space: An illustrative example using *d*-QPSO to generate a locally *D*-optimal approximate design on an irregular design space.
- **S6.** The *d*-QPSO Algorithm for Finding the Optimal Designs for the Odor Removal Experiment: Provides the C++ code allowing the reader to reproduce the results for the odor removal study.

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