

DYNAMIC REGRESSION WITH RECURRENT EVENTS

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RECURRENT EVENTS

Multiple events of the same type for a subject in longitudinal studies. Examples include:

- seizures in epileptic patients;
- successive tumors in cancer patients.

By the multiplicity nature, a dependence structure between events is often observed within a subject; not all dependence is captured by observed covariates, i.e., unobserved heterogeneity between individuals.

THE MOTIVATION FOR DYNAMIC MODEL: CONSTANT EFFECT OVER TIME?

Most models with recurrent events assume constant effects of covariates, e.g., Andersen and Gill (1982) and Lin et al. (2000).

In practice, [the effects may vary over time](#). In a clinical study for AIDS patients, for example,

- a drug may take time to reach its full efficacy,
- the treatment effect may erode over time as drug resistance develops;

e.g., Eshleman et al. (2001) and Wu et al. (2005).

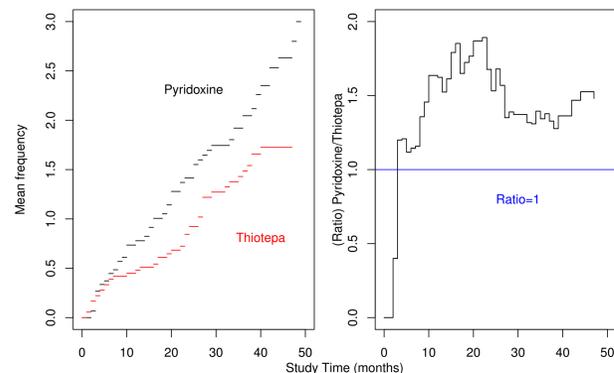


Figure: Nonparametric Nelson-Aalen type mean frequency functions for two treatment arms in the bladder tumor study (Byar (1980)); and the ratio of these over time

THE PROPOSED MODEL

To accommodate time-varying effects, we propose a marginal dynamic regression model to target the mean frequency of recurrent events,

$$E(N^*(t)|Z) = \mu_0(t) \exp\{b_0(t)^\top Z\} = \exp\{\beta_0(t)^\top \tilde{Z}\}, t \geq 0.$$

- $N^*(t)$: the number of events on interval $[0, t]$
- Z : p -dimensional covariate vector
- $\mu_0(t)$: unspecified baseline mean frequency
- $b_0(t)$: p -dimensional [time-varying](#) coefficient

We adopt the conditional independence censoring assumption on incomplete follow-ups, that is,

$$N^*(\cdot) \perp C \mid Z.$$

THE PROPOSED ESTIMATING INTEGRAL EQUATION

Under the proposed model and the independent censorship, it follows that

$$E\left(\tilde{Z}[N(t) - \int_0^t Y(s) d \exp\{\beta_0(s)^\top \tilde{Z}\}]\right) = 0,$$

where $N(t) = N^*(t \wedge C)$ and $Y(t) = I(C \geq t)$.

Therefore we propose an estimating integral equation, for *all* $t \geq 0$,

$$n^{-1} \sum_{i=1}^n \tilde{Z}_i \left[N_i(t) - \int_0^t Y_i(s) d \exp\{\beta(s)^\top \tilde{Z}_i\} \right] = 0.$$

Based on this, $\beta_0(\cdot)$ is [sequentially estimated](#) over ordered observed event times in the sample, cf. Peng and Huang (2007).

LARGE SAMPLE PROPERTIES

THEOREM 1: UNIFORM CONSISTENCY

Under regularity conditions C1–C6, $\sup_{t \in [\kappa, \tau]} \|\hat{\beta}(t) - \beta_0(t)\| \rightarrow 0$, almost surely.

THEOREM 2: WEAK CONVERGENCE

Under regularity conditions C1–C6, $n^{1/2}\{\hat{\beta}(\cdot) - \beta_0(\cdot)\}$ on (κ, τ) weakly converges to a mean-zero Gaussian process.

THE PROPOSED BOOTSTRAP INFERENCE PROCEDURE

For interval estimation, we propose a multiplier bootstrap, adapting Rubin (1981); for *all* $t \geq 0$,

$$\sum_{i=1}^n \xi_i \tilde{Z}_i \left[dN_i(t) - Y_i(t) d \exp\{\beta(t)^\top \tilde{Z}_i\} \right] = 0.$$

- $\{\xi_i\}_{i=1}^n$: size n random sample from $\text{Exp}(1)$
- $\beta^*(t)$: the stochastic solution at time t

The $100(1 - \alpha)\%$ confidence interval for $\beta_0(t)$ can be constructed with the $(\alpha/2)$ th and $(1 - \alpha/2)$ th quantiles of [the empirical distribution for \$\beta^*\(t\)\$](#) .

SIMULATION STUDIES

Monte Carlo simulations with various setups demonstrated the proposed estimator was virtually unbiased and efficient, cp. Fine et al. (2004).

t	$h_0(t) = \log(t)$				$h_0(t) = \exp\{-t/\exp(1)\}$											
	Proposed Method	Fine et al. (2004)			Proposed Method	Fine et al. (2004)										
	B	SD	SE	Cov95	B	SD	SE	Cov95								
Multiplicative unit-mean gamma frailty with variance 1																
0.5	-8	240	234	94.5	-11	257	247	93.9	-6	393	380	95.0	-3	421	401	94.0
1.0	-1	209	203	93.7	2	230	229	94.3	-11	350	335	93.0	-23	388	380	93.8
1.5	0	202	194	93.6	-4	256	243	93.1	-11	339	324	93.3	-22	442	405	93.8
2.0	-5	203	194	93.2	-16	301	278	92.6	-6	342	324	92.8	-8	518	469	92.6
2.5	-9	213	199	93.1	-38	424	370	91.4	0	360	333	92.6	-1	725	625	90.0

NOTE: B: empirical bias ($\times 1000$); SD: empirical standard deviation ($\times 1000$); SE: average standard error ($\times 1000$); Cov95: empirical coverage probability of the Wald 95% confidence interval ($\times 100$). Based on 1,000 Monte Carlo replications.

APPLICATION TO THE BLADDER TUMOR DATA

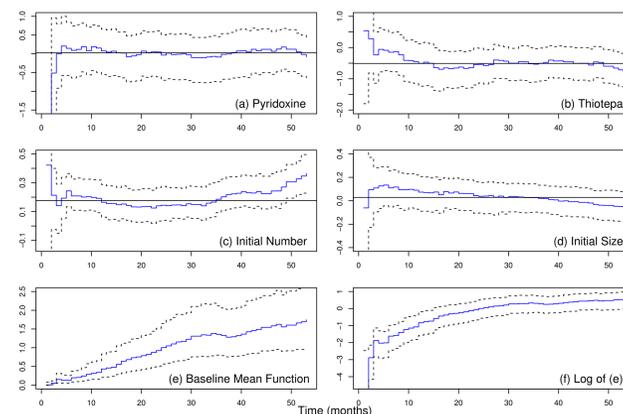


Figure: Estimates for time-varying effects of covariates and the baseline mean frequency function (blue rugged lines); with the point-wise 95% bootstrap percentile confidence intervals (dashed lines).

REMARKS

- We propose a marginal model on mean frequency of recurrent events, accommodating *dynamic* effects of covariates, cf. the *proportional* means model of Lin et al. (2000);
- It is a *global* model over time for evolving effects of covariates, which facilitates efficient estimation, cf. Fine et al. (2004)'s *local* model.
- Consistency and weak convergence of the proposed estimator are established.
- The proposed nonparametric bootstrap inference procedure provides confidence band construction even for an infinite-dimensional quantity.
- Conducted simulations and two real data analyses illustrated practical utility of the proposed method.

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