

Some problems in robustness of design

Doug Wiens

Statistics Centre

Dept. of Mathematical and Statistical Sciences

University of Alberta

- Why robustness of design?
 - The models for which we construct designs are generally at best approximations.
 - The ‘best’ design for a slightly wrong model can be much more than slightly sub-optimal.
 - Although we will fit the assumed, ‘ideal’ model, we should design for protection against biases arising from any of a range (‘neighbourhood’) of alternate models.
- Illustrations to follow are based on simple straight line regression (SLR) – even here there are intriguing, and frustratingly difficult, open problems.
- More complex situations discussed later, as time allows.

- SLR: We fit the model

$$E [Y (x)] = \mathbf{f}' (x) \boldsymbol{\theta} = \theta_0 + \theta_1 x,$$

(i.e. $\mathbf{f}' (x) = (1, x)$) for $x \in [-1, 1]$ but are concerned that

$$E [Y (x)] = \mathbf{f}' (x) \boldsymbol{\theta} + \psi (x),$$

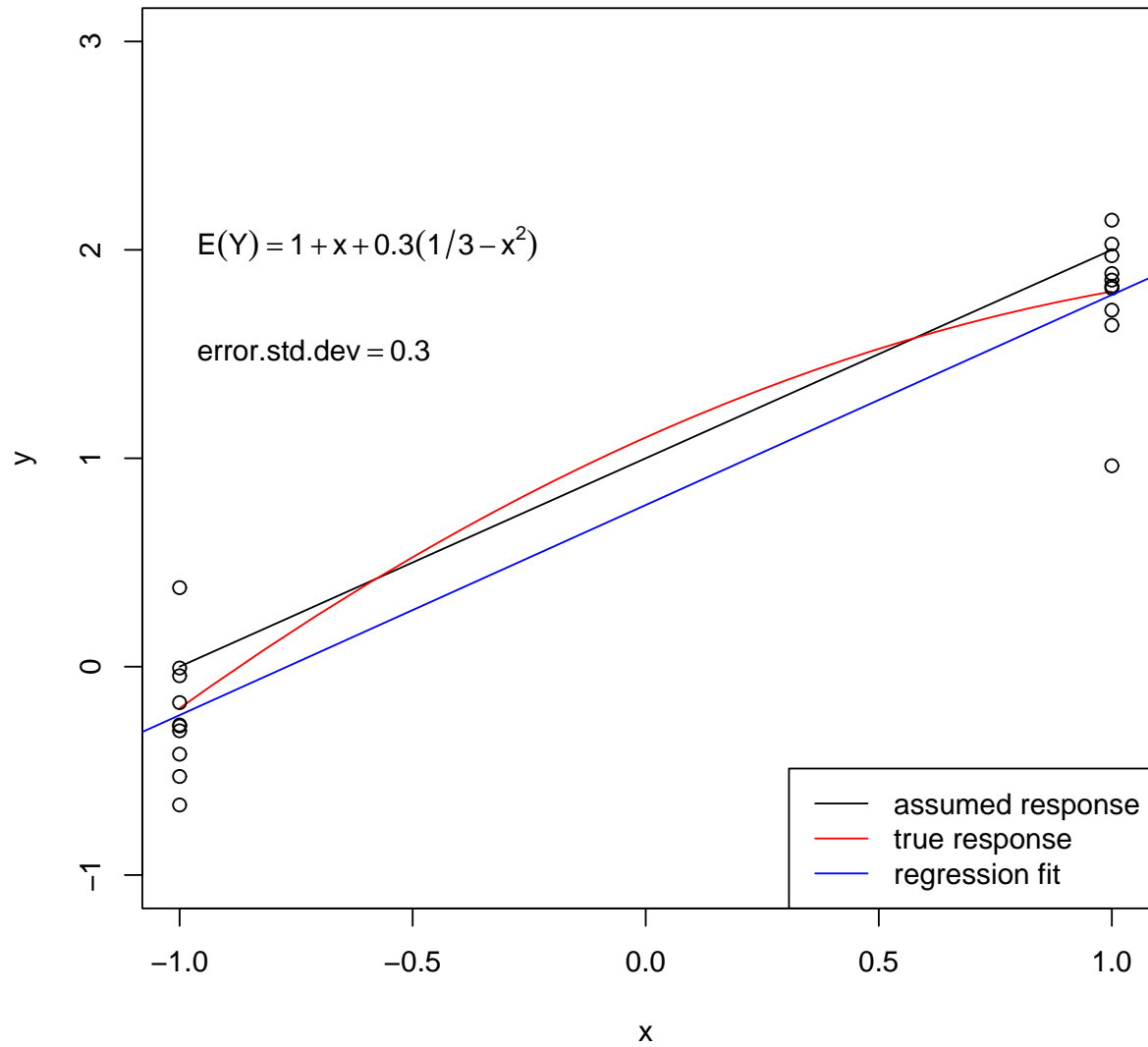
for some 'small' function ψ .

- Example: $\psi (x) = \alpha \cdot (1/3 - x^2)$ for some 'small' α – quadratic 'contamination', orthogonal to \mathbf{f} :

$$\int_{-1}^1 \mathbf{f} (x) \psi (x) dx = \begin{pmatrix} 0 \\ 0 \end{pmatrix}; \quad (1)$$

this can always be arranged and ensures that $\boldsymbol{\theta}$ is well defined.

Classically 'optimal' design with misspecified response



- Let ξ be the ‘design’, sometimes viewed as a pmf $\xi(x_i) = n_i/n$ if n_i observations are to be made at the design point x_i , at other times viewed as the ‘design measure’ or the edf of the design points. For instance the lse $\hat{\boldsymbol{\theta}} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{y}$ has covariance matrix $\sigma_\varepsilon^2 (\mathbf{X}'\mathbf{X})^{-1}$

$$\begin{aligned}
 &= \frac{\sigma_\varepsilon^2}{n} \left(\sum \frac{n_i}{n} \mathbf{f}(x_i) \mathbf{f}'(x_i) \right)^{-1} = \frac{\sigma_\varepsilon^2}{n} \left(\sum \xi(x_i) \mathbf{f}(x_i) \mathbf{f}'(x_i) \right)^{-1} \\
 &= \frac{\sigma_\varepsilon^2}{n} \mathbf{M}_\xi^{-1}, \text{ where } \mathbf{M}_\xi = \int_{-1}^1 \mathbf{f}(x) \mathbf{f}'(x) \xi(dx).
 \end{aligned}$$

- If $E[Y(x)] = \mathbf{f}'(x) \boldsymbol{\theta} + \psi(x)$ then the presence of ψ introduces a bias:

$$E[\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}] = \mathbf{M}_\xi^{-1} \mathbf{b}_{\psi, \xi}, \text{ where } \mathbf{b}_{\psi, \xi} = \int_{-1}^1 \mathbf{f}(x) \psi(x) \xi(dx).$$

- Compare (1) – the continuous uniform ‘design’ is unbiased. [‘Space-filling’]
- The mse matrix of $\hat{\boldsymbol{\theta}}$ is

$$\text{MSE}(\psi, \xi) = \frac{\sigma_\varepsilon^2}{n} \mathbf{M}_\xi^{-1} + \mathbf{M}_\xi^{-1} \mathbf{b}_{\psi, \xi} \mathbf{b}'_{\psi, \xi} \mathbf{M}_\xi^{-1}.$$

- For a function $\phi(\cdot)$ like the determinant ('D-optimality'), trace ('A-optimality'), max eigenvalue ('E-optimality') etc. we might aim to minimize (the worst case of) $\phi(\text{MSE}(\psi, \xi))$:

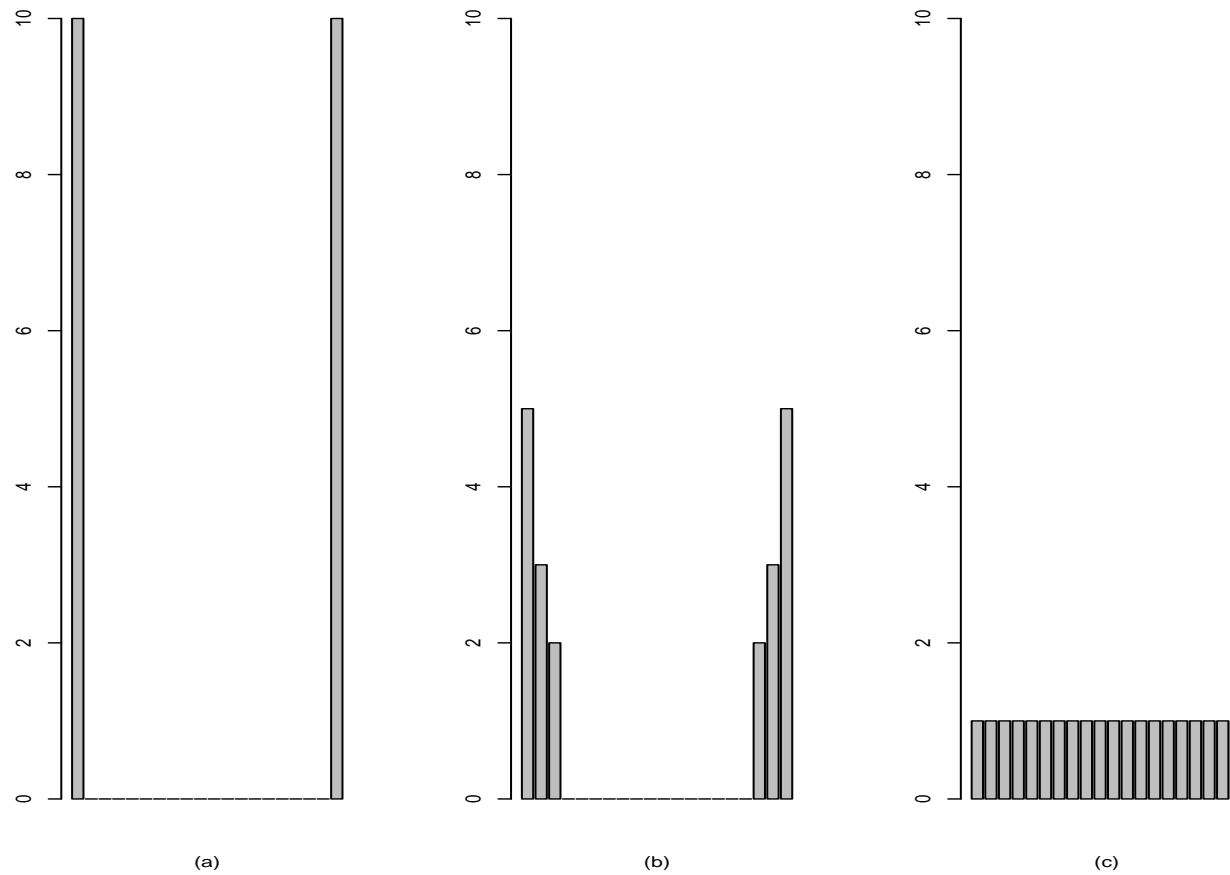
$$\min_{\xi} \max_{\psi} \phi(\text{MSE}(\psi, \xi)).$$

- The max value $\phi(\text{MSE}(\psi_{\max}, \xi))$ turns out to depend only on the maximum eigenvalue of this *mse* matrix (or one closely related to it); these eigenvalues are functions of the design itself.

- Back to SLR - if (as has always been assumed in these problems) the design is symmetric ($\xi(x) = \xi(-x)$) then

$$\mathbf{M}_{\xi} = \int_{-1}^1 \mathbf{f}(x) \mathbf{f}'(x) \xi(dx) = \int_{-1}^1 \begin{pmatrix} 1 & x \\ x & x^2 \end{pmatrix} \xi(dx) = \begin{pmatrix} 1 & 0 \\ 0 & \mu_2(\xi) \end{pmatrix},$$

and it turns out that the other matrices involved in $\text{MSE}(\psi_{\max}, \xi)$ are also diagonal, so that the eigenvalues can be read off; now proceed to minimize the larger (?) of them.



Implementations of minimax designs, $n = 20$; max subject to $\int \psi^2(x)dx \leq \tau^2/n$. (a) $\tau^2 = 0$, (b) τ^2 positive but small, (c) $\tau^2 \rightarrow \infty$.

- **Problem 1:** Does symmetry have to be assumed, or is the optimal ('minimax') design necessarily symmetric?
 - Possible route to a solution: Drop this assumption and determine the optimal design without it. This is complicated (for me) – the various matrices determining the loss, which are diagonal under symmetry, are no longer so.
 - Sometimes, as for D-optimality, there are nice convexity properties – the loss is smaller at the average $(\xi_+ + \xi_-) / 2$ of the design for ' x ' and the design for ' $-x$ ', and this average design is symmetric. Not true for A- or E-optimality (although it is in 'classical' theory).
 - For the 'classical' theory all of these loss function are convex and symmetry is indeed necessary.

- **Problem 2:** For D-optimality, the optimal design is range-invariant: if the regression range $[-1, 1]$ is changed to $[a, b]$ then we linearly transform each design point x_i and put the same number of observations at the transformed design point as there were at x_i . This is not so for (classical or robust) A- and E-optimality – the design weights will also change. How do they vary with the range?
 - A solution to Problem 2 might shed light on the next problem – perhaps, in the case of SLR, it is a problem only for certain ranges $[a, b]$ and not for others.

- **Problem 3:** The minimax problem for SLR described above was reduced to finding an optimal design ξ_0 minimizing the maximum of two quantities:

$$\max(\lambda_1(\xi), \lambda_2(\xi)).$$

- What are these quantities? First (recall $\int \psi^2(x)dx \leq \tau^2/n$) define

$$\nu = \frac{\tau^2}{\tau^2 + \sigma_\varepsilon^2} \in [0, 1],$$

so $\nu = 0 \Rightarrow$ only variance of interest, $\nu = 1 \Rightarrow$ only bias of interest. Put $\mu_2 = \int x^2 \xi(dx)$. Let $m(x)$ be the density of the 'design' ξ (!) – the optimal design has to be approximated in the end, to implement it. Then – for instance – in the case of E-optimality and the regression range $\chi = [-1/2, 1/2]$ we have

$$\lambda_1(\xi) = 1 - \nu + \nu \left[\int_{\chi} (m(x) - 1)^2 dx \right],$$

$$\lambda_2(\xi) = \frac{1 - \nu}{\mu_2} + \nu \left[\int_{\chi} x^2 \left(\frac{m(x)}{\mu_2} - 12 \right)^2 dx \right].$$

- We are to minimize $\max(\lambda_1(\xi), \lambda_2(\xi))$. A relatively easy route to a solution – when it works – is:

1. Hope that λ_1 will be the largest one, find ξ_1 minimizing $\lambda_1(\xi)$, verify that $\max(\lambda_1(\xi_1), \lambda_2(\xi_1)) = \lambda_1(\xi_1)$.
2. If not, do it with λ_2 : find ξ_2 minimizing $\lambda_2(\xi)$, verify that $\max(\lambda_1(\xi_2), \lambda_2(\xi_2)) = \lambda_2(\xi_2)$.

BUT this doesn't always work; for A- and E-optimality, often

$$\begin{aligned}\max(\lambda_1(\xi_1), \lambda_2(\xi_1)) &= \lambda_2(\xi_1) \text{ and} \\ \max(\lambda_1(\xi_2), \lambda_2(\xi_2)) &= \lambda_1(\xi_2).\end{aligned}$$

- Shi, Ye, Zhou (2003) – non-smooth optimization methods (maximum eigenvalues aren't 'smooth') – obtain a description of the solution; severe computational problems.
- Alternate approach, being attempted for SLR:
 1. Find ξ_1 minimizing $\lambda_1(\xi)$ in the class for which $\lambda_1(\xi) > \lambda_2(\xi)$.
 2. Find ξ_2 minimizing $\lambda_2(\xi)$ in the class for which $\lambda_2(\xi) > \lambda_1(\xi)$.

Choose the one with the smaller maximum loss: put

$$\xi_0 = \begin{cases} \xi_1, & \text{if } \lambda_1(\xi_1) < \lambda_2(\xi_2), \\ \xi_2, & \text{otherwise.} \end{cases}$$

- Even if this works out for SLR it remains a problem for more complex models – approximate quadratic regression is the ‘next’ one.
- Problem 3 is not related to the non-invariance of A- and E-optimality under changes of the regression range. For quadratic regression it also arises in the case of D-optimality, even though this criterion is range-invariant.
- Sometimes the best ‘solution’ is to avoid the problem altogether:
 - Restrict the class of allowable designs - e.g. tractable parametric classes with the parameters optimally chosen.
 - Finite design space? – Avoids the problem of designs with densities (designs are now equivalent to vectors); integer optimization problems.

- A sampling of other problems in robustness of design:
 - Spatial sampling – look at robustness of the design against the ‘wrong’ spatial correlation function (as well as against the wrong response function); applications in computer experiments too.
 - Survey sampling - survey samplers resist model-based designs, but could perhaps be convinced otherwise if the models were sufficiently flexible.
 - Response surface exploration – sequentially search for a maximum on a surface, allowing for model errors.
 - Threshold designs – search for a level of the inputs at which the response exceeds a particular threshold (in the face of model uncertainty).

There is no shortage of problems to be 'robustified'.